Noise covariance matrix estimation with subspace model identification for Kalman filtering

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Noise covariance matrix estimation

Outline

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MOTIVATIONS AND FRAMEWORK

PhD thesis with Michelin

- Future autonomous vehicles will require well-developed Advanced Driver Assistance Systems (ADAS) to assist human beings in driving.
- One path chosen by Michelin for ADAS improvement consists in providing the ADAS with information related to the state of the road.
- Such information is included in the grip potential quantity.
- Benefits for passenger security (to name a few) are
 - detection of roads with low-grip area,
 - evaluation of the driving conditions,
 - reduction of the impact of rear end collisions.



Problem formulation



• The grip potential is

$$\mu_{\rm max} = \max{\left(\frac{\sqrt{F_x^2+F_y^2}}{F_z}\right)}, \label{eq:max}$$

i.e., the maximum effort a tire can generate before sliding on the road.

Problem formulation (cont'd)

Problem

Estimate the grip potential under standard driving conditions from sensors fitted on production vehicles.



Problem formulation (cont'd)

- Getting (noisy) data requires to measure the friction μ and the slip ratio s.
- No dedicated sensors exist on production vehicles.
- These signals must be estimated knowing that, for the longitudinal dynamics,

$$\mu = \frac{F_x}{F_z},$$

$$s = \frac{\omega R_{\rm rol} - v_x}{\max\left(\omega R_{\rm rol} - v_x\right)}.$$

 A Kalman filter is suggested to reconstruct the components of μ and s accurately.

KALMAN FILTERING: A REMINDER

Toy example

 Let us assume we want¹ to determine from remote noisy measurements the position and speed (state) of a cart moving straightforward.



¹See Understanding the basis of the Kalman filter via a simple and minimized intuitive derivation, R. Faragher, IEEE Signal Processing Magazine, 2012.

- We need
 - a model of the cart dynamics,
 - a model of the measuring process,
 - a description of the noise and uncertainties acting on the system.
- Because we get new measurements every $T_{\!s\!}\text{,}$ most of the Kalman filters
 - are based on dynamical systems discretized in the time domain,
 - can be updated recursively (only the estimated state from the previous time step and the current measurement are needed to compute the estimate for the current state).
- Beside estimating the state vector recursively, the Kalman filters propagate and update its uncertainty as soon as new noisy measurements are available.

















Discrete time linear Kalman filter

• The standard discrete time linear Kalman filter considers models evolving as follows

$$egin{aligned} oldsymbol{x}_{k+1} &= oldsymbol{F}_k oldsymbol{x}_k + oldsymbol{G}_k oldsymbol{u}_k + oldsymbol{w}_k, \ oldsymbol{y}_k &= oldsymbol{H}_k oldsymbol{x}_k + oldsymbol{v}_k, \end{aligned}$$

where, for $k \in \mathbb{T} = \{0, \cdots, N-1\}$,

- $v_k \in \mathbb{R}^{n_y imes 1}$ stands for a time sample of the observation or output measurement noise sequence $(\mathbf{v}_k)_{k \in \mathbb{Z}}$,
- $\boldsymbol{w}_k \in \mathbb{R}^{n_x imes 1}$ stands for a time sample of the process noise sequence $(\boldsymbol{\mathsf{w}}_k)_{k \in \mathbb{Z}}$,
- The sequences $(m{v}_i)_{i\in\mathbb{T}}$ and $(m{w}_i)_{i\in\mathbb{T}}$ are used to describe
 - the noise acting on the real system,
 - the (in)accuracy of the model representation,
 - the confidence we have in the model and the measurements.

Discrete time linear Kalman filter

• Let us consider a system, the behavior of which is governed by the vector difference equation

$$\mathbf{x}_{k+1} = \boldsymbol{F}_k \mathbf{x}_k + \boldsymbol{G}_k \boldsymbol{u}_k + \mathbf{w}_k,$$

where F_k , G_k and u_k are assumed to be perfectly known while the initial state and its symmetric positive definite covariance are defined as follows

$$\mathbb{E} \left\{ \mathbf{x}_0 \right\} = \mathbf{x}_0,$$
$$\mathbb{E} \left\{ \left(\mathbf{x}_0 - \mathbf{x}_0 \right) \left(\mathbf{x}_0 - \mathbf{x}_0 \right)^\top \right\} = \mathbf{X}_0 \succ 0.$$



• Let us assume that the process disturbances can be described by a zero mean white noise with a finite and symmetric covariance matrix satisfying

$$\mathbb{E} \{ \mathbf{w}_k \} = \mathbf{0},$$
$$\mathbb{E} \{ \mathbf{w}_k \mathbf{w}_j^\top \} = \mathbf{W}_k \delta_{kj}, \ \mathbf{W}_k \succ 0.$$

• Then, the uncorrected state and error covariance matrix propagate from the previous corrected estimates as follows

$$egin{aligned} \hat{\mathbf{x}}_k^- &= oldsymbol{F}_{k-1} \hat{\mathbf{x}}_{k-1}^+ + oldsymbol{G}_{k-1} oldsymbol{u}_{k-1}, \ oldsymbol{X}_k^- &= oldsymbol{F}_{k-1} oldsymbol{X}_{k-1}^+ oldsymbol{F}_{k-1}^ op + oldsymbol{W}_{k-1}, \end{aligned}$$

where the notations $^-$ and $^+$ stand for "before and after new measurements".

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Once a new measurement is available, *i.e.*, once we measure

$$\mathbf{y}_k = \boldsymbol{H}_k \mathbf{x}_k + \mathbf{v}_k,$$

where H_k is known and by assuming that the output disturbances are also described by a zero mean white noise (with finite and symmetric covariance matrix) uncorrelated with the process noise, *i.e.*,

$$\begin{split} \mathbb{E} \left\{ \mathbf{v}_k \right\} &= \mathbf{0}, \\ \mathbb{E} \left\{ \mathbf{v}_k \mathbf{v}_j^\top \right\} &= \mathbf{V}_k \delta_{kj}, \ \mathbf{V}_k \succ 0, \\ \mathbb{E} \left\{ \mathbf{v}_k \mathbf{w}_j^\top \right\} &= \mathbf{0}, \text{ for all } k \text{ and } j, \end{split}$$

we can first update the Kalman gain

$$oldsymbol{K}_k = oldsymbol{X}_k^- oldsymbol{H}_k^ op oldsymbol{\left(oldsymbol{H}_k oldsymbol{X}_k^- oldsymbol{H}_k^ op + oldsymbol{V}_k
ight)^{-1}.$$

• The state estimate is updated as follows

$$\hat{\mathbf{x}}_k^+ = \hat{\mathbf{x}}_k^- + \mathbf{K}_k(\mathbf{y}_k - \mathbf{H}_k \hat{\mathbf{x}}_k^-),$$

while the error covariance estimate update satisfies

$$oldsymbol{X}_k^+ = \left(oldsymbol{I}_{n_x imes n_x} - oldsymbol{K}_k oldsymbol{H}_k
ight)oldsymbol{X}_k^-$$
 .



• The state estimate is updated as follows

$$\hat{\mathbf{x}}_k^+ = \hat{\mathbf{x}}_k^- + \mathbf{K}_k(\mathbf{y}_k - \mathbf{H}_k \hat{\mathbf{x}}_k^-),$$

while the error covariance estimate update satisfies

$$oldsymbol{X}_k^+ = \left(oldsymbol{I}_{n_x imes n_x} - oldsymbol{K}_k oldsymbol{H}_k
ight)oldsymbol{X}_k^-.$$

- In order to get reliable results, the Kalman filter requires
 - input and output measurements,
 - matrices F_k, G_k and H_k generated from (a discretization and linearization of) the physical equations governing the system behavior,
 - values for V_k and W_k , $k \in \mathbb{T}$.

Time invariant conditions

In the sequel, for $k \in \mathbb{T}$,

$$F_k = F$$
, $G_k = G$, $H_k = H$, $V_k = V$, $W_k = W$.

 Because the matrices V and W are used to describe the confidence we have in the model and the measurements, we aim at determining them by comparing

• the model used in the Kalman filter,

- a model estimated from the available data sets.
- Herein, the data driven model learning solution is a subspace based model identification approach.



Notations interlude...

• For any vector $r_k \in \mathbb{R}^{n_r \times 1}$ and parameters M, i and $\ell \in \mathbb{N}^+_*$, we define

$$oldsymbol{r}_{i,M} = egin{bmatrix} oldsymbol{r}_i \ oldsymbol{r}_{i+1} \ dots \ oldsymbol{r}_{i+M-1} \end{bmatrix} \in \mathbb{R}^{Mn_r imes 1},$$
 $oldsymbol{R}_{i,M} = egin{bmatrix} oldsymbol{r}_i & oldsymbol{r}_{i+1} \cdots & oldsymbol{r}_{i+M-1} \end{bmatrix} \in \mathbb{R}^{n_r imes M},$

and the block Hankel matrix as follows

$$oldsymbol{R}_{i,\ell,M} = egin{bmatrix} oldsymbol{r}_i & oldsymbol{r}_{i+1} & oldsymbol{r}_{i+2} & \cdots & oldsymbol{r}_{i+M} \ oldsymbol{r}_{i+\ell-1} & oldsymbol{r}_{i+\ell} & \cdots & oldsymbol{r}_{i+M+\ell-2} \end{bmatrix} \in \mathbb{R}^{\ell n_r imes M}.$$

Basic idea of our solution (cont'd)

- The subspace based model learning methods yield estimates of $\boldsymbol{x}_t, t \in \{f, \dots, N-1\} \times T_s$, and $\{\boldsymbol{F}, \boldsymbol{G}, \boldsymbol{H}\}$ up to a similarity transformation.
- We can thus estimate²

$$egin{bmatrix} \hat{m{W}}_{f,M-1} \ \hat{m{V}}_{f,M-1} \end{bmatrix} = egin{bmatrix} \hat{m{X}}_{f+1,M} \ m{Y}_{f,M-1} \end{bmatrix} - egin{bmatrix} m{F} & m{G} \ m{H} & m{0} \end{bmatrix} egin{bmatrix} \hat{m{X}}_{f,M-1} \ m{U}_{f,M-1} \end{bmatrix},$$

and

$$\begin{bmatrix} \hat{\boldsymbol{V}} & \hat{\boldsymbol{S}} \\ \hat{\boldsymbol{S}}^{\top} & \hat{\boldsymbol{W}} \end{bmatrix} = \lim_{M \to \infty} \frac{1}{M} \begin{bmatrix} \hat{\boldsymbol{W}}_{f,M-1} \\ \hat{\boldsymbol{V}}_{f,M-1} \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{W}}_{f,M-1}^{\top} & \hat{\boldsymbol{V}}_{f,M-1}^{\top} \end{bmatrix}$$

²We have
$$M = N - f$$

SUBSPACE STATE SPACE MODEL IDENTIFICATION: A REMINDER

Problem formulation

 Subspace model learning (SML) methods mainly focus on state space models of finite dimensional linear time invariant causal and discrete time dynamical systems described by

$$oldsymbol{x}_{k+1} = oldsymbol{A} oldsymbol{x}_k + oldsymbol{B} oldsymbol{u}_k + oldsymbol{w}_k,$$
 (1a)

$$oldsymbol{y}_k = oldsymbol{C} oldsymbol{x}_k + oldsymbol{v}_k.$$
 (1b)

where $k \in \mathbb{T}$ whereas the noise sources are assumed to be realizations of zero mean white noise statically independent of the input sequence whereas

$$\mathbb{E}\left[\begin{bmatrix}\mathbf{v}_i\\\mathbf{w}_i\end{bmatrix}\begin{bmatrix}\mathbf{v}_j^\top & \mathbf{w}_j^\top\end{bmatrix}\right] = \begin{bmatrix}\mathbf{V} & \mathbf{S}\\\mathbf{S}^\top & \mathbf{W}\end{bmatrix}\delta_{ij}.$$
 (2)

Problem formulation (cont'd)

- By assuming that
 - A1 the input vector sequence is quasi stationary and exciting of sufficient order,
 - A2 the pair (A, C) is observable and the pair $(A, \begin{bmatrix} B & V^{1/2} \end{bmatrix})$ is reachable,
 - standard SML solutions aim at estimating consistently
 - the order n_x of the system,
 - $({m A},{m B},{m C})$ up to a similarity transformation,
 - K such that an approximated minimum variance estimate of x_k, k ∈ T, can be reconstructed.



Innovation form

Under Assumption A2, Kalman filtering theory proves that, for every choice of (v_i)_{i∈Z} and (w_i)_{i∈Z} satisfying Eq. (2), a matrix K ∈ ℝ^{n_y×1} an a zero mean white noise sequence (e_i)_{i∈Z} exist such that the innovation form

$$egin{aligned} \hat{oldsymbol{x}}_{k+1} &= oldsymbol{A}\hat{oldsymbol{x}}_k + oldsymbol{B}oldsymbol{u}_k + oldsymbol{K}oldsymbol{e}_k, \ oldsymbol{y}_k &= oldsymbol{C}\hat{oldsymbol{x}}_k + oldsymbol{e}_k, \end{aligned}$$

is equivalent to Eq. (1) (same I/O behavior).

• Furthermore, the matrix K satisfies

$$\lambda_{\max}(\boldsymbol{A} - \boldsymbol{K}\boldsymbol{C}) < 1.$$

 Standard SML solutions estimate (A, B, C, K) by using this innovation form.

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Notations interlude... (cont'd)

 With matrices A, B, C and D of appropriate dimensions, for ℓ ≥ n_x, the extended controllability matrix can be defined as follows

$$\Omega_\ell(oldsymbol{A},oldsymbol{B}) = egin{bmatrix} oldsymbol{A}^{\ell-1}oldsymbol{B} & \cdots & oldsymbol{A}oldsymbol{B} & oldsymbol{B} \end{bmatrix}.$$

• We also define the extended observability matrix

$$oldsymbol{\Gamma}_\ell(oldsymbol{A},oldsymbol{C}) = egin{bmatrix} oldsymbol{C}^ op & (oldsymbol{C}oldsymbol{A})^ op & (oldsymbol{C}oldsymbol{A}^{\ell-1})^ op \end{bmatrix}^ op,$$

and the block lower triangular Toeplitz matrix

$$oldsymbol{\Delta}_\ell(oldsymbol{A},oldsymbol{B},oldsymbol{C},oldsymbol{D}) = egin{bmatrix} oldsymbol{D} & oldsymbol{0} & o$$



First recursions

• With f and $\ell \in \mathbb{N}^*$,

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$$egin{aligned} oldsymbol{y}_\ell &= oldsymbol{C} \hat{oldsymbol{x}}_\ell + oldsymbol{e}_\ell, \ oldsymbol{y}_{\ell+1} &= oldsymbol{C} oldsymbol{A} \hat{oldsymbol{x}}_\ell + oldsymbol{C} oldsymbol{B} oldsymbol{u}_\ell + oldsymbol{C} oldsymbol{K} oldsymbol{e}_\ell + oldsymbol{e}_{\ell+1}, \ oldsymbol{y}_{\ell+2} &= \cdots, \end{aligned}$$

i.e.,

$$m{y}_{\ell,f} = \Gamma_f(m{A},m{C}) \hat{m{x}}_\ell + \Delta^u_f(m{A},m{B},m{C},m{0})m{u}_{\ell,f} + m{n}_{\ell,f},$$
 with

$$oldsymbol{n}_{\ell,f} = oldsymbol{\Delta}_f^e(oldsymbol{A},oldsymbol{K},oldsymbol{C},oldsymbol{I}_{n_y imes n_y})oldsymbol{e}_{\ell,f}.$$



First recursions (cont'd)

By using all the available time shifts, we get

$$oldsymbol{Y}_{\ell,f,M} = oldsymbol{\Gamma}_f(oldsymbol{A},oldsymbol{C}) \hat{oldsymbol{X}}_{\ell,M} + oldsymbol{\Delta}_f^u oldsymbol{U}_{\ell,f,M} + oldsymbol{N}_{\ell,f,M},$$

where $M=N+1-\ell-f$ whereas

$$egin{aligned} oldsymbol{\Delta}_f^u &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{B},oldsymbol{C},oldsymbol{0}), \ oldsymbol{N}_{\ell,f,M} &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{K},oldsymbol{C},oldsymbol{I},oldsymbol{M},oldsymbol{L}_{\ell,f,M}) \ oldsymbol{\Delta}_f^e &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{B},oldsymbol{C},oldsymbol{0}), \ oldsymbol{N}_{\ell,f,M} &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{K},oldsymbol{C},oldsymbol{I},oldsymbol{M},oldsymbol{L}_{\ell,f,M}) \ oldsymbol{E}_{\ell,f,M} &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{K},oldsymbol{C},oldsymbol{I},oldsymbol{M},oldsymbol{L}_{\ell,f,M}) \ oldsymbol{E}_{\ell,f,M} &= oldsymbol{\Delta}_f(oldsymbol{A},oldsymbol{K},oldsymbol{C},oldsymbol{L}_{\ell,f,M}) \ oldsymbol{E}_{\ell,f,M}. \end{aligned}$$



Predictor state space form

• Because, for $k \in \mathbb{T}$,

$$\boldsymbol{e}_k = \boldsymbol{y}_k - \boldsymbol{C}\hat{\boldsymbol{x}}_k,$$

we get for $k\in\mathbb{T}$

$$egin{aligned} \hat{m{x}}_{k+1} &= ilde{m{A}}\hat{m{x}}_k + m{B}m{u}_k + m{K}m{y}_k, \ m{y}_k &= m{C}\hat{m{x}}_k + m{e}_k, \end{aligned}$$

with

$$\tilde{A} = A - KC.$$



Further recursions

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• We can easily see that

$$egin{aligned} \hat{oldsymbol{x}}_\ell &= ilde{oldsymbol{A}} \hat{oldsymbol{x}}_{\ell-1} + oldsymbol{B} oldsymbol{u}_{\ell-1} + oldsymbol{K} oldsymbol{y}_{\ell-1}, \ &= ilde{oldsymbol{A}}^2 \hat{oldsymbol{x}}_{\ell-2} + ilde{oldsymbol{A}} oldsymbol{B} oldsymbol{u}_{\ell-2} + oldsymbol{A} oldsymbol{K} oldsymbol{y}_{\ell-2} + oldsymbol{B} oldsymbol{u}_{\ell-1} + oldsymbol{K} oldsymbol{y}_{\ell-1}, \end{aligned}$$

i.e, for
$$p \in \mathbb{N}^*$$

 $\hat{x}_{\ell} = \tilde{A}^p \hat{x}_{\ell-p} + \Omega_p(\tilde{A}, K) y_{\ell-p,p} + \Omega_p(\tilde{A}, B) u_{\ell-p,p}.$
Because $\lambda_{\max} \left(A - KC \right) < 1$,
 $\bar{x}_{\ell} = \Omega_p(\tilde{A}, K) y_{\ell-p,p} + \Omega_p(\tilde{A}, B) u_{\ell-p,p},$

is the the optimal linear estimate of \hat{x}_ℓ (in the mean square error sense) given $y_{\ell-p,p}$ and $u_{\ell-p,p}$.

Data equation

• Thus, with $\ell \ge p$,

$$\hat{oldsymbol{X}}_{\ell,M}pproxar{oldsymbol{X}}_{\ell,M}=igg[oldsymbol{\Omega}_p(ilde{oldsymbol{A}},oldsymbol{B}) \quad oldsymbol{\Omega}_p(ilde{oldsymbol{A}},oldsymbol{K})igg]igg[oldsymbol{U}_{\ell-p,p,\ell+M-1}\ oldsymbol{Y}_{\ell-p,p,\ell+M-1}igg]$$

• A straightforward combination of the former equations lead to the data equation

$$egin{aligned} \mathbf{Y}_{\ell,f,M} &= oldsymbol{\Delta}_{f}^{u} oldsymbol{U}_{f,\ell,M} + oldsymbol{N}_{f,\ell,M} + oldsymbol{\Gamma}_{f}(oldsymbol{A},oldsymbol{C}) \ & imes \left[oldsymbol{\Omega}_{p}(ilde{oldsymbol{A}},oldsymbol{B}) \quad oldsymbol{\Omega}_{p}(ilde{oldsymbol{A}},oldsymbol{K})
ight] egin{bmatrix} oldsymbol{U}_{\ell-p,p,\ell+M-1} \ oldsymbol{Y}_{\ell-p,p,\ell+M-1} \end{bmatrix} \end{aligned}$$

 This is nothing but a standard linear least squares minimization problem with good statistical properties because the past I/O data are uncorrelated with the future noise.

Numerical implementation

 As any linear least squares problem, an LQ factorization is involved. More precisely,

$$egin{bmatrix} egin{aligned} egin{aligne} egin{aligned} egin{aligned} egin{aligned} egin$$

It can be proved that

$$\lim_{N\to\infty} \boldsymbol{L}_{32} \boldsymbol{L}_{22}^{-1} \begin{bmatrix} \boldsymbol{U}_p \\ \boldsymbol{Y}_p \end{bmatrix} = \boldsymbol{\Gamma}_f(\boldsymbol{A}, \boldsymbol{C}) \hat{\boldsymbol{X}}_{f,M}.$$



Numerical implementation (cont'd)

Via the following SVD,

$$\boldsymbol{L}_{32}\boldsymbol{L}_{22}^{-1}\begin{bmatrix} \boldsymbol{U}_p\\ \boldsymbol{Y}_p \end{bmatrix} = \boldsymbol{\mathcal{U}}\boldsymbol{\Sigma}\boldsymbol{\mathcal{V}}^{\top},$$

we get an estimate of the system order n_x and

$$\hat{\boldsymbol{\Gamma}}_{f}(\boldsymbol{A},\boldsymbol{C}) = \boldsymbol{\mathcal{U}}(:,1:\hat{n}_{x})\boldsymbol{\Sigma}^{1/2}(1:\hat{n}_{x},1:\hat{n}_{x}),$$
$$\hat{\hat{\boldsymbol{X}}}_{f,M} = \boldsymbol{\Sigma}^{1/2}(1:\hat{n}_{x},1:\hat{n}_{x})\boldsymbol{\mathcal{V}}^{\top}(1:\hat{n}_{x},:).$$

• Reliable estimates of (A, B, C, K) can be generated from $\hat{\Gamma}_f(A, C)$ and $\hat{\hat{X}}_{f,M}$ straightforwardly (up to a similarity transformation T).

NOISE COVARIANCE MATRIX ESTIMATION

Similarity transformation

- For our noise covariance matrix estimation problem, we need to have $\hat{\hat{X}}_{f,M}$ in the basis of (F, G, H).
- Fortunately, knowing $\hat{\Gamma}_f(m{A},m{C})$ and $\Gamma_f(m{F},m{G})$, we have

$$\Gamma_f(\boldsymbol{F}, \boldsymbol{G})\boldsymbol{T} = \hat{\Gamma}_f(\boldsymbol{A}, \boldsymbol{C}).$$

- Thus, we can get $\hat{\hat{X}}_{f,M}$ is the correct basis as follows

$$\begin{split} \hat{\hat{\boldsymbol{X}}}_{f,M_T} &= \boldsymbol{\Gamma}_f^{\dagger}(\boldsymbol{F},\boldsymbol{G})\hat{\boldsymbol{\Gamma}}_f(\boldsymbol{A},\boldsymbol{C}) \\ &\times \boldsymbol{\Sigma}^{1/2}(1:\hat{n}_x,1:\hat{n}_x)\boldsymbol{\mathcal{V}}^{\top}(1:\hat{n}_x,:). \end{split}$$



Covariance matrix estimation

• We can finally estimate

$$egin{bmatrix} \hat{m{W}}_{f,M-1} \ \hat{m{V}}_{f,M-1} \end{bmatrix} = egin{bmatrix} \hat{m{X}}_{f+1,M_T} \ m{Y}_{f,M-1} \end{bmatrix} - egin{bmatrix} m{F} & m{G} \ m{H} & m{0} \end{bmatrix} egin{bmatrix} \hat{m{X}}_{f,M-1_T} \ m{U}_{f,M-1} \end{bmatrix},$$

and

$$\begin{bmatrix} \hat{\boldsymbol{V}} & \hat{\boldsymbol{S}} \\ \hat{\boldsymbol{S}}^{\top} & \hat{\boldsymbol{W}} \end{bmatrix} = \lim_{M \to \infty} \frac{1}{M} \begin{bmatrix} \hat{\boldsymbol{W}}_{f,M-1} \\ \hat{\boldsymbol{V}}_{f,M-1} \end{bmatrix} \begin{bmatrix} \hat{\boldsymbol{W}}_{f,M-1}^{\top} & \hat{\boldsymbol{V}}_{f,M-1}^{\top} \end{bmatrix}.$$



NUMERICAL ILLUSTATIONS • We consider

$$\begin{split} \boldsymbol{F} &= \begin{bmatrix} 0.603 & 0.603 & 0 & 0 \\ -0.603 & 0.603, 0 & 0 \\ 0 & 0 & -0.603 & -0.603 \\ 0 & 0 & 0.603 & -0.603 \end{bmatrix}, \\ \boldsymbol{G} &= \begin{bmatrix} 1.1650 & -0.6965 \\ 0.6268 & 1.6961 \\ 0.0751 & 0.0591 \\ 0.3516 & 1.7971 \end{bmatrix}, \\ \boldsymbol{H} &= \begin{bmatrix} 0.2641 & -1.4462 & 1.2460 & 0.5774 \\ 0.8717 & -0.7012 & -0.6390 & -0.3600 \\ 0.8717 & -0.7012 & -0.6390 & -0.3600 \end{bmatrix}, \\ \boldsymbol{K} &= 4 \times \begin{bmatrix} 0.1242 & -0.0895 \\ -0.0828 & -0.0128 \\ 0.0390 & -0.0968 \\ -0.0225 & 0.1459 \end{bmatrix}, \\ \boldsymbol{R}_e &= \begin{bmatrix} 0.0176 & -0.0267 \\ -0.0267 & 0.0497 \end{bmatrix}. \end{split}$$

• We generate 10^3 realizations of the noise sequence and we select a data length N = 1000.

Estimates of the elements of $oldsymbol{V}$

		\hat{v}_{11}	\hat{v}_{12}	\hat{v}_{22}
Theo. value		0.0176	-0.0267	0.0497
Sample cov.	avg.	0.0176	-0.0267	0.0497
	std.	0.0008	0.0012	0.0022
ICM	avg.	0.0588	-0.0750	0.0888
	std.	0.0412	0.0149	0.0057
DCM	avg.	0.031	-0.041	0.028
	std.	0.0041	0.0013	0.006
CMM	avg.	0.023	-0.073	0.031
	std.	0.0087	0.0066	0.0092
New meth.	avg.	0.0198	-0.0272	0.0516
	std.	0.0011	0.0013	0.0024



Estimates of the elements of $oldsymbol{W}$

		\hat{w}_{11}	\hat{w}_{12}	\hat{w}_{22}	\hat{w}_{23}	\hat{w}_{34}	\hat{w}_{44}
Theo. value		0.0202	-0.0045	0.0149	-0.0198	0.0012	-0.0031
Sample cov.	avg.	0.0202	-0.0045	0.0149	-0.0198	0.0012	-0.0031
	std.	0.8886e-03	0.2054e-03	0.6563e-03	0.8750e-03	0.0509e-03	0.1487e-03
ICM	avg.	0.0526	-0.0150	0.0355	-0.0454	0.0041	-0.0103
	std.	0.0146	0.0079	0.0066	0.0067	0.0042	0.0036
DCM	avg.	0.0113	-0.0058	0.0186	-0.0285	0.003	-0.0103
	std.	0.004	0.0033	0.0068	0.0067	0.0039	0.0033
CMM	avg.	0.0170	-0.0041	0.0124	-0.0234	0.0041	-0.0043
	std.	0.0097	0.0082	0.0064	0.0062	0.0052	0.0028
New meth.	avg.	0.0196	-0.0041	0.0145	-0.0190	0.0015	-0.0026
	std.	0.0017	0.0006	0.0011	0.0011	0.0004	0.0005



Estimates of the elements of old S

		\hat{s}_{11}	\hat{s}_{13}	\hat{s}_{21}	\hat{s}_{24}
Theo. value		0.0183	-0.0045	0.0131	-0.0172
Sample cov.	avg.	0.0183	-0.0045	0.0131	-0.0172
	std.	0.0008	0.0002	0.0006	0.0008
New meth.	avg.	0.0181	-0.0039	0.0137	-0.0169
	std.	0.0011	0.0007	0.0010	0.0010



Michelin results

- Data is generated with VI-CRT (realistic simulator).
- A nonlinear state space model is used, *i.e.*,

$$\dot{\boldsymbol{x}}(t) = \boldsymbol{f}(\boldsymbol{x}(t), \boldsymbol{u}(t), t, \boldsymbol{\theta}), \\ \boldsymbol{y}(t) = \boldsymbol{g}(\boldsymbol{x}(t), t, \boldsymbol{\theta}),$$

with

$$\boldsymbol{x} = \begin{bmatrix} v_x & \omega_f & \omega_r & F_{x_f} & F_{x_r} & \dot{F}_{x_f} & \dot{F}_{x_r} & \kappa & \dot{\kappa} \end{bmatrix}^\top, \\ \boldsymbol{u} = \begin{bmatrix} T_f & T_r \end{bmatrix}^\top, \\ \boldsymbol{y} = \begin{bmatrix} v_x & \omega_f & \omega_r & \dot{\kappa} \end{bmatrix}^\top,$$

involving

- a single track model for the dynamics,
- an effective tire radius model,
- a suspension model and a load transfer model.







Measurements used by the observer, SNR = 25 dB



Long. tire force estimation





Vehicle speed estimation





Friction reconstruction



Université "Poitiers

Slip ratio reconstruction





CONCLUSIONS AND DISCUSSION

- A new solution for noise covariance matrix estimation has been introduced.
- This solution involves SML solutions only.
- No tuning is required by the user.
- The accuracy of this solution was proved in the linear Kalman filter framework.
- This solution can be used with an EKF (no accuracy proof yet).
- See https://doi.org/10.1002/acs.3213.
- Do not hesitate to use SML solutions.

